Lecture 04: Modelling Sequences

DT2118 Speech and Speaker Recognition

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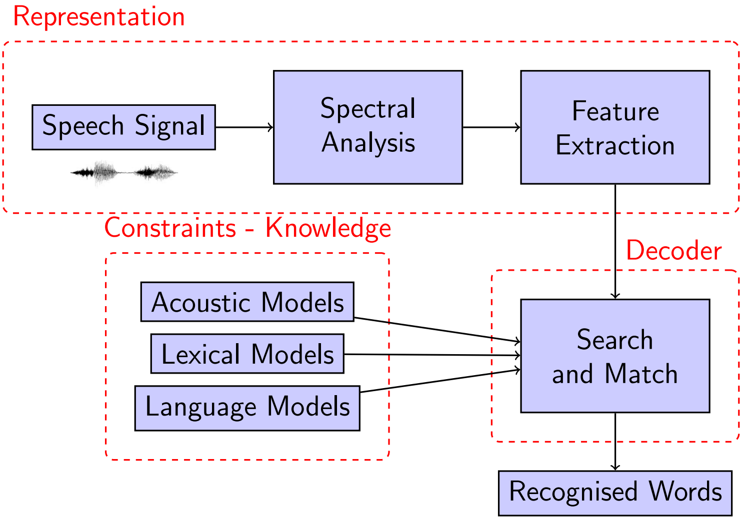
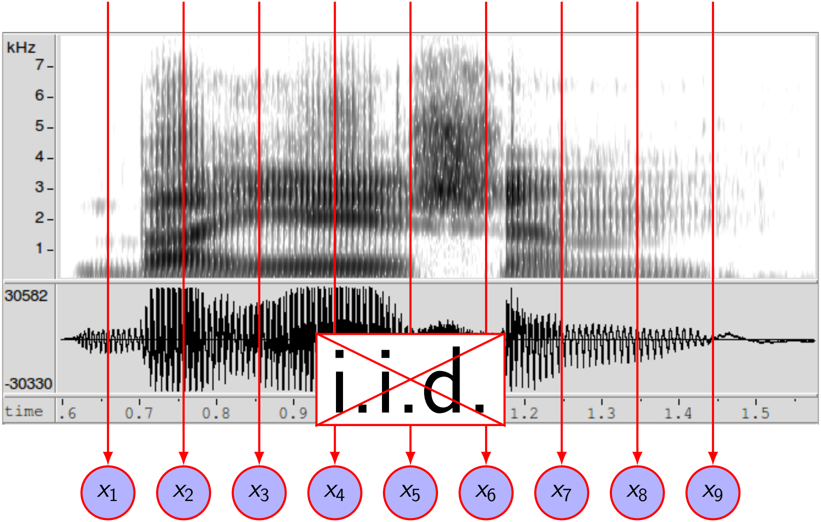
Components of ASR System



Sequences in Statistical Terms



Sequential Data: Not Only Speech

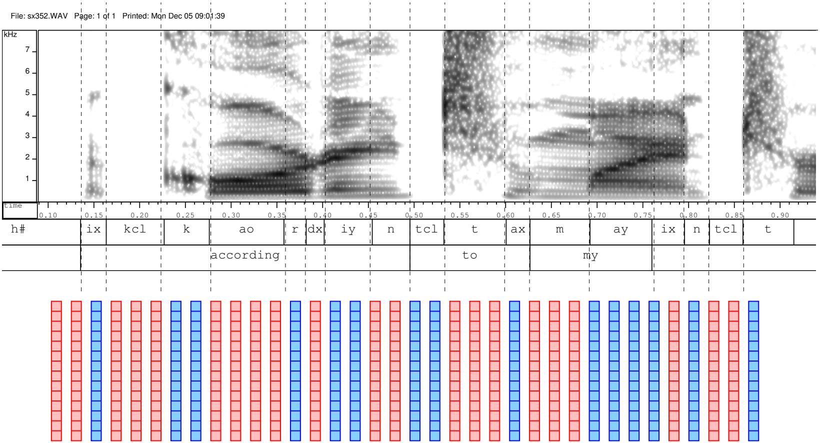
Time sequences

Timeless sequences

text

Nel mezzo del cammin di nostra vita
mi ritrovai per una selva oscura,
ché la diritta via era smarrita.
Ahi quanto a dir qual era è cosa dura
esta selva selvaggia e aspra e forte
che nel pensier rinova la paura!
Tant' è amara che poco è più morte;
ma per trattar del ben ch'i' vi trovai,
dirò de l'altre cose ch'i' v'ho scorte.
Io non so ben ridir com' i' v'intrai,
tant' era pien di sonno a quel punto
che la verace via abbandonai.

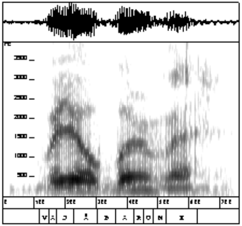
Frame-Based Processing



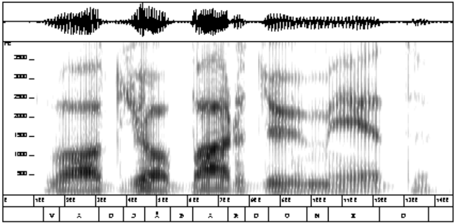
Frame-wise distance metrics

distance	$d(x, y)$
city block:	$\sum_i \lvert x_i - y_i \rvert$
Euclidean:	$\sqrt{\sum_i (x_i - y_i)^2}$
Mahalanobis:	$\sum_i (x_i - \mu_y)^2 / \sigma_y$
probability function:	$f(X = x \lvert \mu_y, \Sigma_y)$
artificial neural networks:	$f\left(\sum_i w_i x_i - \theta\right)$

Comparing Utterances



Va jobbaru me



Vad jobbar du med

"What is your occupation"
("What work you with")

Combining frame-wise scores into utterance scores

Template Matching

- oldest technique
- simple comparison of template patterns
- compensate for varying speech rate (Dynamic Programming)

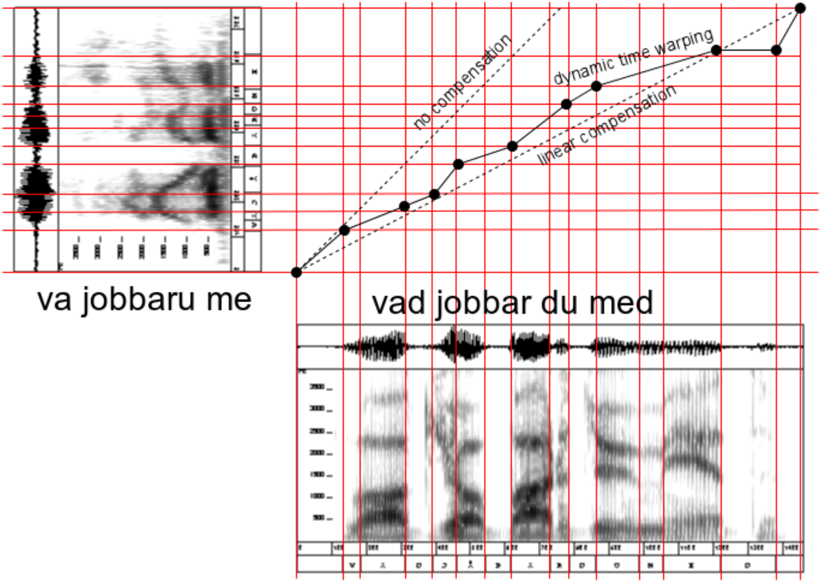
Hidden Markov Models (HMMs)

- most used technique
- models of segmental structure of speech
- recognition by Viterbi search (Dynamic Programming)

Template Matching: Why?

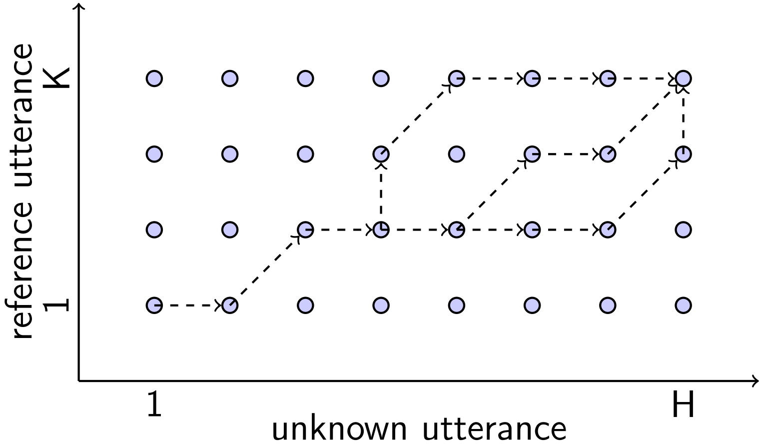
1. Historical: first technique used in ASR
2. Pedagogical: explain the problem and Dynamic Programming
3. Research: it is getting “hot” again

Template Matching



Dynamic Programming

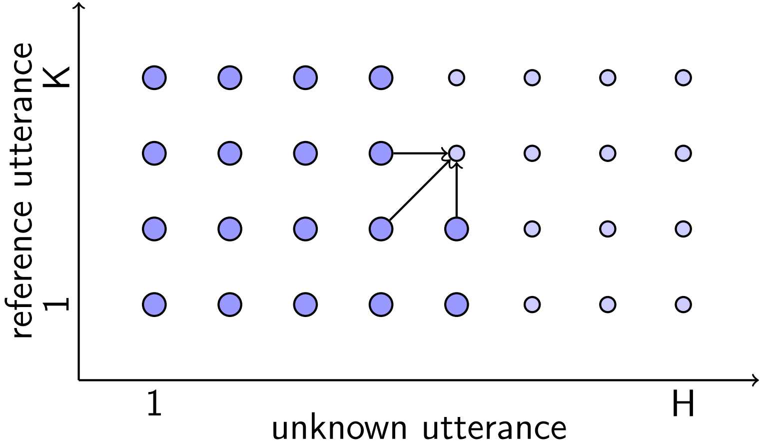
- compare any possible alignment
- problem: exponential with H and K!



Dynamic Programming

Dynamic Time Warping (DTW) algorithm

for $h = 1$ to H **do**
    for $k = 1$ to K **do**
        $AccD[h,k] = LocD[h,k] + \min(AccD[h-1,k], AccD[h-1,k-1], AccD[h,k-1])$



DP Example: Spelling

- observations are letters
- local distance: 0 (same letter), 1 (different letter)
- Unknown utterance: ALLDRIG
- Reference1: ALDRIG
- Reference2: ALLTID
- Problem: find closest match

Distance char-by-char:

- ALLDRIG–ALDRIG $= 5$
- ALLDRIG–ALLTID $= 4$

DP Example: Solution

LocD[h,k]=

G	1	1	1	1	1	1	0
I	1	1	1	1	1	0	1
R	1	1	1	1	0	1	1
D	1	1	1	0	1	1	1
L	1	0	0	1	1	1	1
A	0	1	1	1	1	1	1
	A	L	L	D	R	I	G

AccD[h,k]=

G	5	4	4	3	2	1	0
I	4	3	3	2	1	0	1
R	3	2	2	1	0	1	2
D	2	1	1	0	1	2	3
L	1	0	0	1	2	3	4
A	0	1	2	3	4	5	6
	A	L	L	D	R	I	G

Distance ALLDRIG–ALDRIG: $AccD[\mathrm{H,K}] = 0$

Distance ALLDRIG–ALLTID? (5min)

DP Example: Solution

$LocD$[h,k]=

D	1	1	1	0	1	1	1
I	1	1	1	1	1	0	1
T	1	1	1	1	1	1	1
L	1	0	0	1	1	1	1
L	1	0	0	1	1	1	1
A	0	1	1	1	1	1	1
	A	L	L	D	R	I	G

$AccD$[h,k]=

D	5	3	3	2	3	3	3
I	4	2	2	2	2	2	3
T	3	1	1	1	2	3	4
L	2	0	0	1	2	3	4
L	1	0	0	1	2	3	4
A	0	1	2	3	4	5	6
	A	L	L	D	R	I	G

Distance ALLDRIG–ALDRIG: $AccD[H,K] = 0$

Distance ALLDRIG–ALLTID: $AccD[H,K] = 3$

Best path: Backtracking

Sometimes we want to know the path

1. at each point [h,k] remember the minimum distance predecessor (back pointer)
2. at the end point [H,K] follow the back pointers until the start

Properties of Template Matching

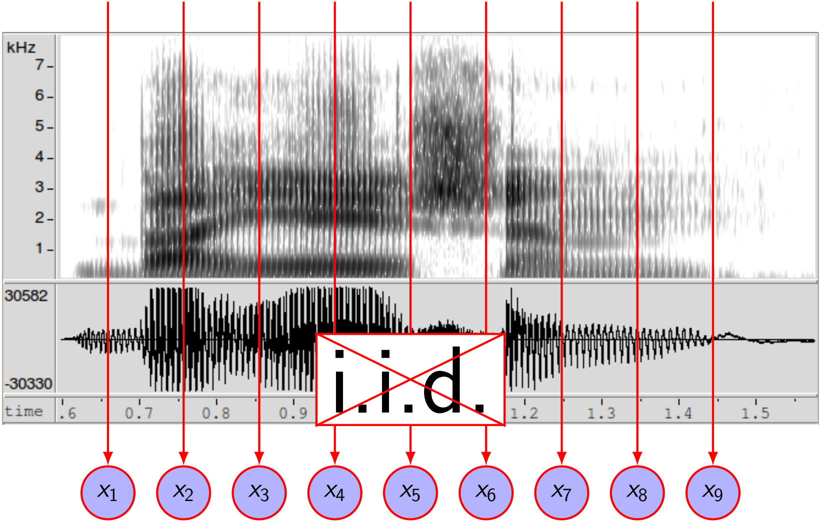
Pros:

+ No need for phonetic transcriptions
+ within-word co-articulation for free
+ high time resolution

Cons:

– cross-word co-articulation not modelled
– requires recordings of every word
– not easy to model variation
– does not scale up with vocabulary size

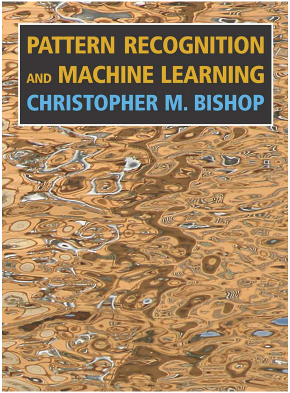
Statistical Approach



Historical Perspective

- Hidden Markov Models first studied in the '60s
- applied to ASR in the end of the '80s
- later seen as special case of Bayesian Networks

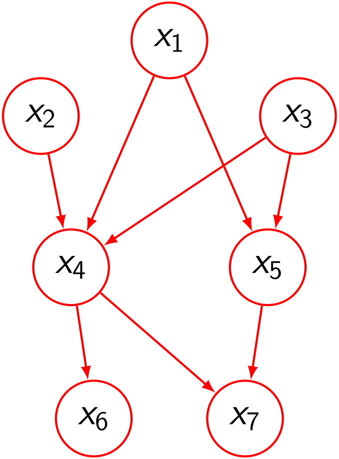
Extra Literature (Optional)



C. M. Bishop. *Pattern Recognition and Machine Learning*. Springer, 2006

Bayesian Networks (reminder)

$$
\begin{aligned}
p(x_1, \ldots, x_7) = \\
&p(x_1) \\
&p(x_2) \\
&p(x_3) \\
&p(x_4 | x_1, x_2, x_3) \\
&p(x_5 | x_1, x_3) \\
&p(x_6 | x_4) \\
&p(x_7 | x_4, x_5)
\end{aligned}
$$

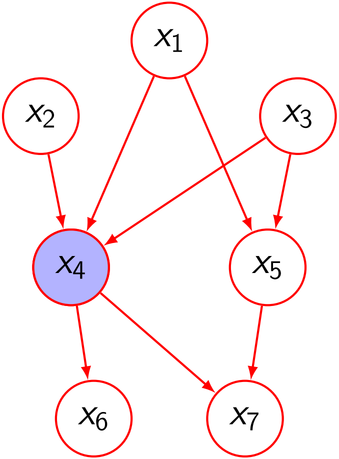


Bayesian Networks (reminder)

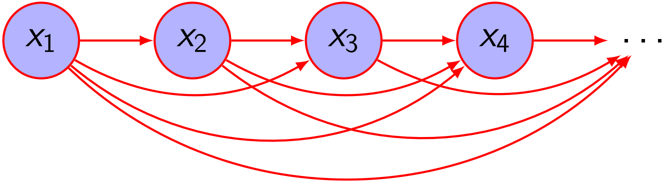
If we observe x_4...

d-separation:

- x_6 and x_7 cond. indep.
- x_1, x_2 and x_3 dependent (head-to-head)



(Dynamic) Bayesian Networks

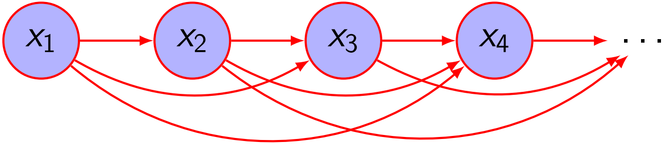


independence assumption (e.g. i.i.d) not satisfactory Most general case, applying product rule recursively ($p(a, b) = p(a)p(b|a)$)

$$p(x_1, \ldots, x_N) = p(x_1)p(x_2|x_1)p(x_3|x_1, x_2) \cdots \\ \cdots p(x_N|x_1, \ldots, x_{N-1})$$

Grows quadratically with sequence length (N)!!!

Markov assumption



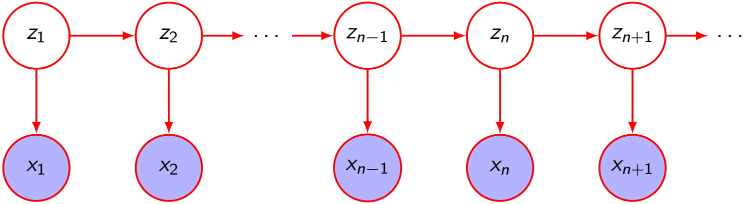
Third order Markov assumption:

$$\begin{aligned} p(x_1, \ldots, x_N) &= p(x_1)p(x_2|x_1)p(x_3|x_1, x_2) \\ &\quad \prod_{n=4}^{N} p(x_n|x_{n-3}, x_{n-2}, x_{n-1}) \end{aligned}$$

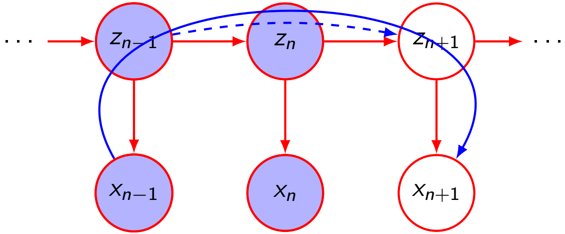
Grows quadratically with order!!!

State Space Models

Adding latent variables z_n



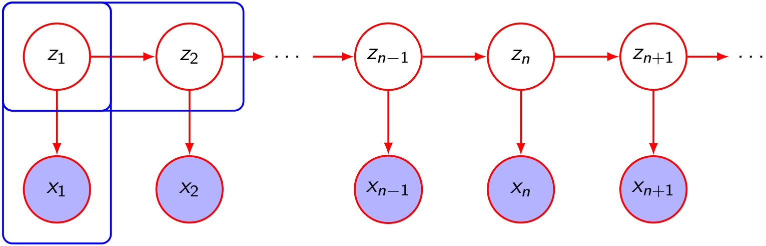
State Space Models: Properties



- given z_n, z_{n+1} is independent of $z_1, \ldots, z_{n-1}$
  $p(z_{n+1}|z_1, \ldots, z_n) = p(z_{n+1}|z_n)$
- $p(x_{n+1}|x_1, \ldots, x_n)$ does not simplify

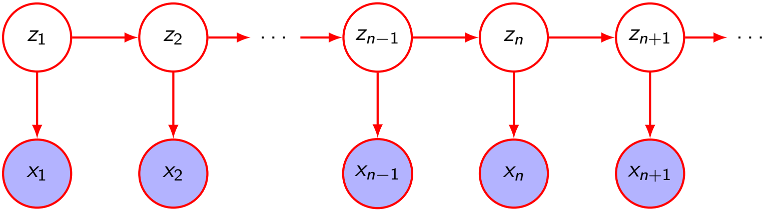
We have modelled indefinitely long dependencies with a limited set of parameters!

Stationary State Space Models



- Emission: $p(x_n|z_n)$
- Transition: $p(z_n|z_{n-1})$
- Initial: $p(z_1)$

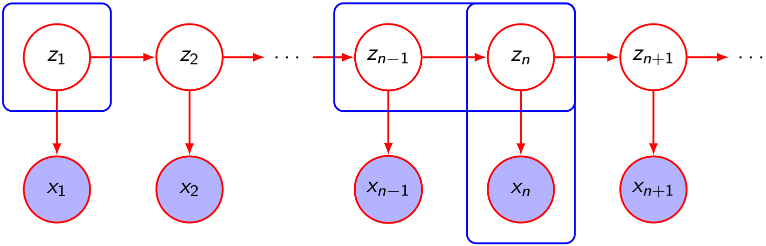
State Space Models Instances



- if z_n are discrete: Hidden Markov Models
- if z_n are continuous: Linear Dynamical Systems

Hidden Markov Models

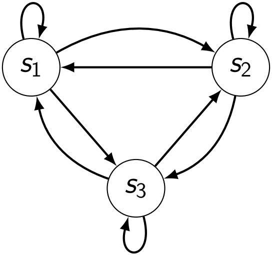
State space models with discrete z_n



- Emission: $p(x_n|z_n) = p(x_n|z_n, \phi)$
  equivalent to Mixture Model
- Transition: $p(z_n|z_{n-1}) = p(z_n|z_{n-1}, A)$
- Initial: $p(z_1) = p(z_1|\pi)$

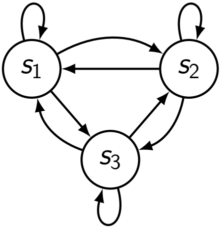
Hidden Markov Models (HMMs)

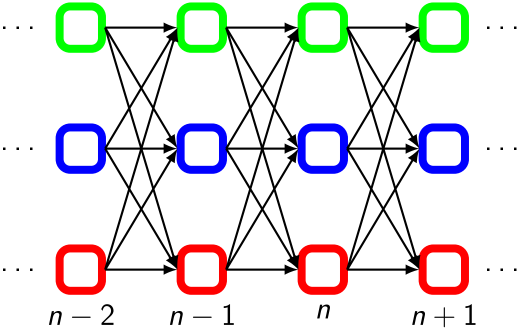
Ergodic HMM



Elements:	
set of states:	$S = \{s_1, s_2, s_3\}$
transition probabilities:	$A(s_a, s_b) = P(s_b, t \mid s_a, t-1)$
prior probabilities:	$\pi(s_a) = P(s_a, t_0)$
state to observation probs:	$\phi(o, s_a) = P(o \mid s_a)$

HMMs: Trellis (Lattice)





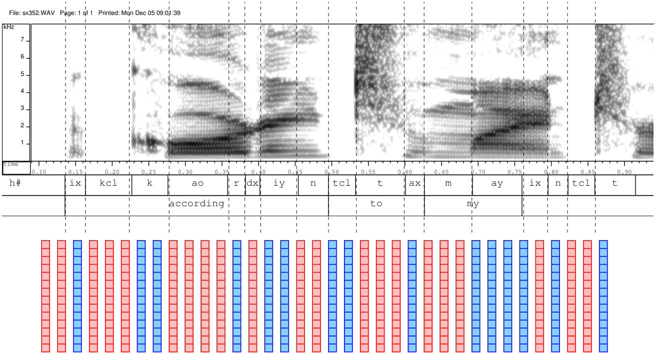
A probabilistic perspective: Bayes' rule

$$P(\text{words}|\text{sounds}) = \frac{P(\text{sounds}|\text{words})P(\text{words})}{P(\text{sounds})}$$

- $P(\text{sounds}|\text{words})$ can be estimated from training data and transcriptions
- $P(\text{words})$: *a priori* probability of the words (Language Model)
- $P(\text{sounds})$: *a priori* probability of the sounds (constant, can be ignored)

Probabilistic Modelling

Problem: How do we model $P(\text{sounds}|\text{words})$?

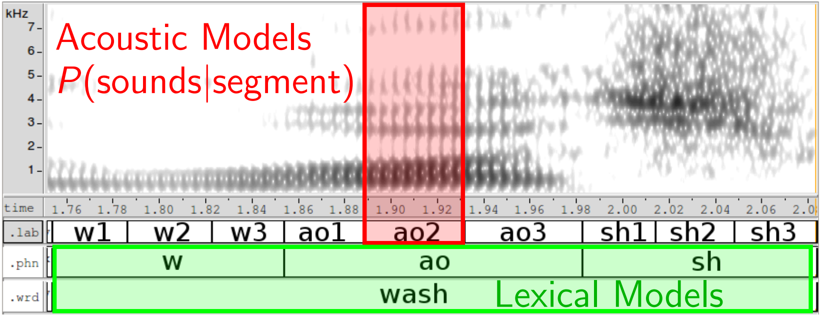


Every feature vector (observation at time t) is a continuous stochastic variable (e.g. MFCC)

Stationarity

Problem: speech is not stationary

- we need to model short segments independently
- the fundamental unit can not be the word, but must be shorter
- usually we model three segments for each phoneme



Local probabilities (frame-wise)

If segment sufficiently short

$$P(\text{sounds}|\text{segment})$$

can be modelled with standard probability distributions

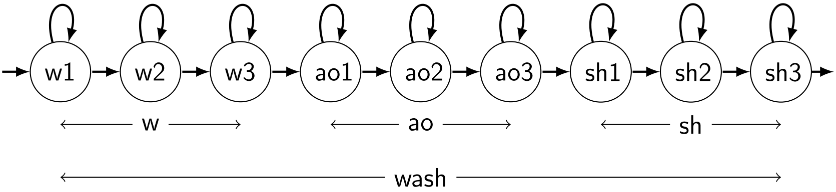
$$\phi(o, s_a) = P(o|s_a)$$

Usually Gaussian or Gaussian Mixture

Global Probabilities (utterance)

Problem: How do we combine the different $P(\text{sounds}|\text{segment})$ to form $P(\text{sounds}|\text{words})$?

Answer: Hidden Markov Model (HMM)



HMM-questions (Inference)

1. what is the probability that the model has generated the sequence of observations? (isolated word recognition) forward algorithm
2. what is the most likely state sequence given the observation sequence? (continuous speech recognition) Viterbi algorithm [3]
3. how can the model parameters be estimated from examples? (training) Baum-Welch[1]

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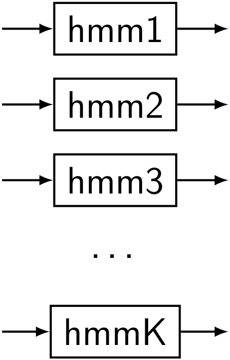
[3] A. J. Viterbi. "Error Bounds for Convolutional Codes and an Asymtotically optimum decoding algorithm". In: *IEEE Trans. Inform. Theory* IT-13 (Apr. 1967), pp. 260–269

[1] L. E. Baum, T. Petrie, G. Soules, and N. Weiss. "A maximization technique occurring in the statistical analysis of probabilistic functions of Markov chains". In: *Ann. Math. Statist.* 41.1 (1970), pp. 164–171

Isolated Words Recognition

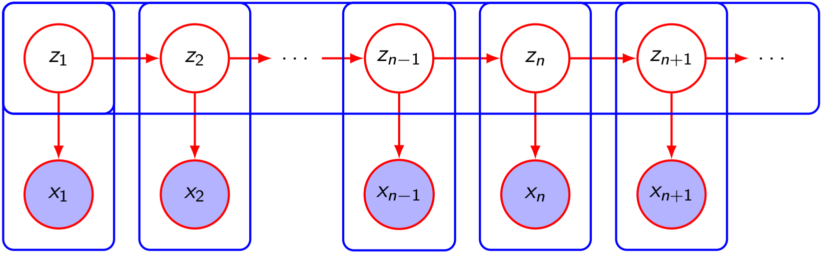


Compare Likelihoods (forward algorithm)

HMM Inference: Joint Distribution

$$X = \{x_1, \ldots, x_N\}$$
$$Z = \{z_1, \ldots, z_N\}$$

$$P(X, Z|\theta) = p(z_1|\pi) \left[\prod_{n=2}^{N} p(z_n|z_{n-1}, A) \right] \prod_{m=1}^{N} p(x_m|z_m, \phi)$$

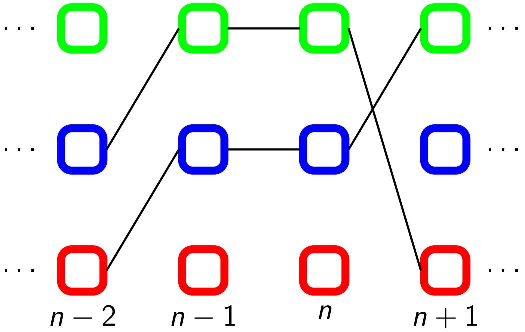


HMM Inference: Likelihood Function

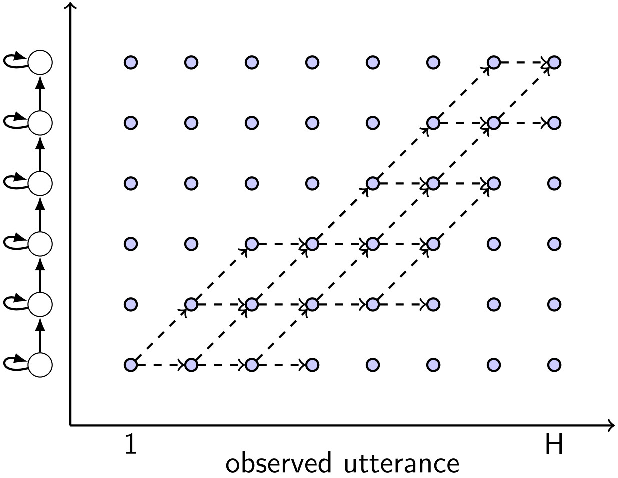
marginalise joint distribution over Z:

$$P(X|\theta) = \sum_Z p(X, Z|\theta)$$

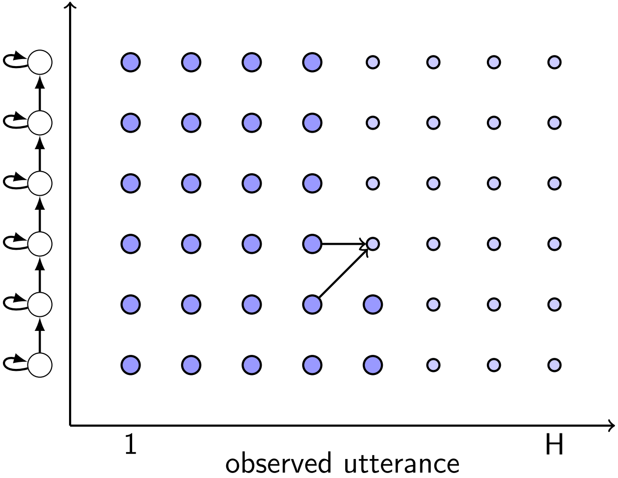
Problem: there are K^N possible sequences for Z



Hidden Markov Models (HMMs)



Hidden Markov Models (HMMs)



Solution: Forward algorithm

Instead of AccD[h,k] (Template Matching)

$$\alpha_n(j) \equiv p(x_1, \ldots, x_n, z_n = s_j)$$

At the end, instead of AccD[H,K]:

$$P(X|\theta) = \sum_{i=1}^{M} \alpha_N(i)$$

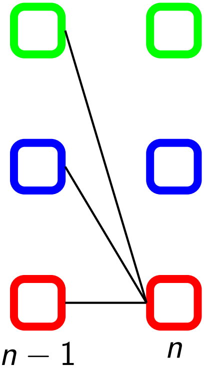
Forward Probability

Initialization:

$$\alpha_1(j) = \pi_j \phi_j(x_1)$$

Recursion:

$$\alpha_n(j) = \left[\sum_{i=1}^{M} \alpha_{n-1}(i) a_{ij} \right] \phi_j(x_n)$$



equivalent to sum-product in Bayesian Networks

Backward probability

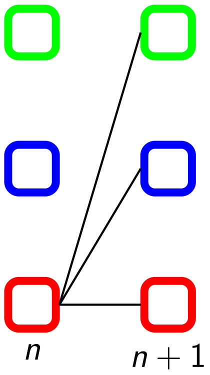
$$\beta_n(i) \equiv p(x_{n+1}, \ldots, x_N | z_n = s_i)$$

Initialization:

$$\beta_N(i) \equiv p(?|z_n = s_i) \equiv 1$$

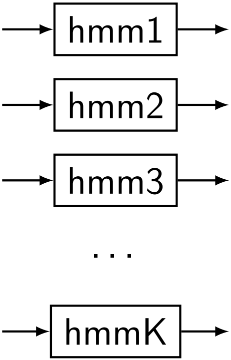
Recursion:

$$\beta_n(i) = \left[\sum_{j=1}^{M} a_{ij} \phi_j(x_{n+1}) \beta_{n+1}(j) \right]$$



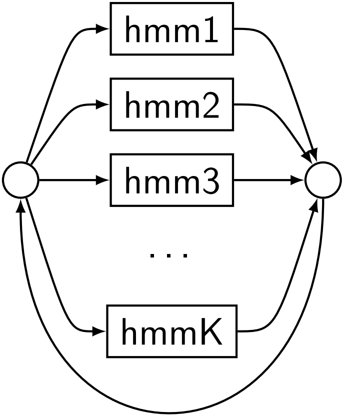
Find best sequence of states: why?

Isolated Words



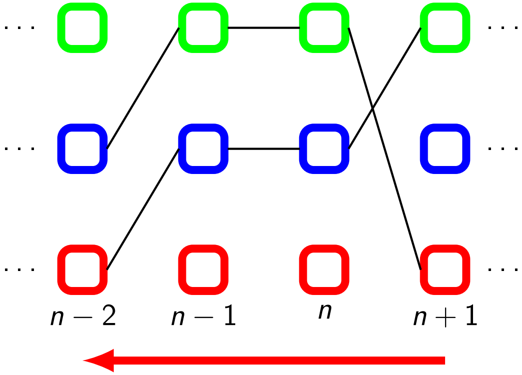
(Likelihood)

Continuous Speech



Find best sequence of states: how?

- Viterbi algorithm [3]
- equivalent to max-sum in Bayesian Networks



[3] A. J. Viterbi. "Error Bounds for Convolutional Codes and an Asymtotically optimum decoding algorithm". In: *IEEE Trans. Inform. Theory* IT-13 (Apr. 1967), pp. 260–269

Summary: update rules

Forward algorithm (sum-product):

$$\alpha_n(j) = \left[\sum_{i=1}^{M} \alpha_{n-1}(i) a_{ij} \right] \phi_j(x_n)$$

Viterbi algorithm (max-sum):

$$V_n(j) = \max_{i=1}^{M} \left[V_{n-1}(i) a_{ij} \right] \phi_j(x_n)$$

$$B_n(j) = \arg\max_{i=1}^{M} \left[V_{n-1}(i) a_{ij} \right]$$

Practical Issues

Product of many probabilities: numerical problems
Solution: work in log domain

$$
\begin{aligned}
\alpha'_1(j) &= \pi'_j + \phi'_j(x_1) \\
\alpha'_n(j) &= \log \sum_{i=1}^{M} e^{\left(\alpha'_{n-1}(i) + a'_{ij}\right)} + \phi'_j(x_n) \\
\beta'_N(i) &= 0 \\
\beta'_n(i) &= \log \sum_{j=1}^{M} e^{\left(a'_{ij} + \phi'_j(x_{n+1}) + \beta'_{n+1}(j)\right)}
\end{aligned}
$$

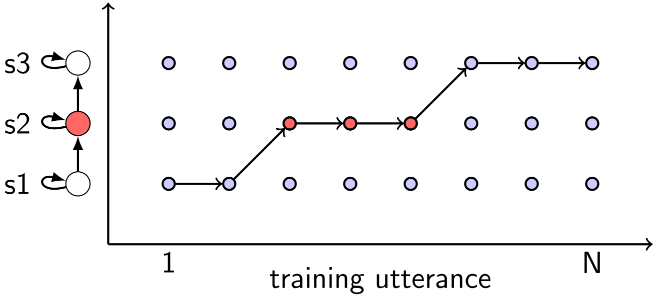
HMM Inference: Learning

- Given observations X update model parameters

$$\theta = \{\pi, A, \phi\}$$

- to maximise either:
  - model fit to data (e.g. likelihood, posterior)
  - classification performance (discriminative training)
- Incomplete data: state sequence Z

Viterbi training (simple approach)



problem: sensitive to misalignments

HMM Inference: Learning

Latent variables $\rightarrow$ Expectation Maximisation

- locally maximise the likelihood
- close form and efficient solution with froward-backward or Baum-Welch algorithm [1]
- general idea: sum over all possible paths weighted by probability of the path
- also: every observation vector contributes to all parameter updates

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[1] L. E. Baum, T. Petrie, G. Soules, and N. Weiss. "A maximization technique occurring in the statistical analysis of probabilistic functions of Markov chains". In: *Ann. Math. Statist.* 41.1 (1970), pp. 164–171

Example: Transition Probability

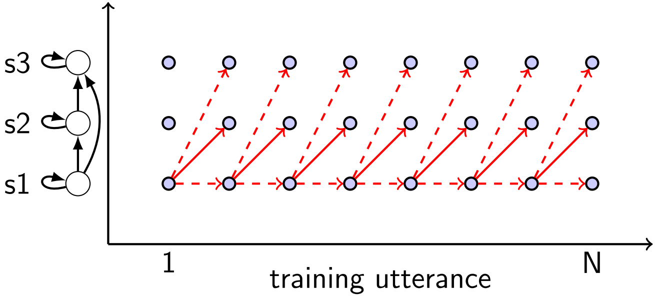
We want to update

$$a_{ij} = P(z_n = s_j | z_{n-1} = s_i)$$

Define $\gamma_n(i, j)$ as the probability of making a transition from s_i to s_j at time step n given the current model parameters θ and the current observation sequence X

$$\gamma_n(i, j) = P(z_{n-1} = s_i, z_n = s_j | X, \theta)$$

Example: Transition Probability



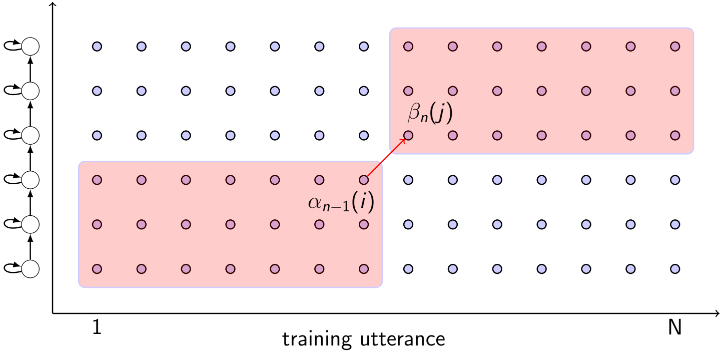
$$a_{12}^{\text{new}} = \frac{E\left[s_1 \to s_2 | X, \theta\right]}{E\left[s_1 \to s_{\text{any}} | X, \theta\right]} = \frac{\sum_{n=1}^{N} \gamma_n(1,2)}{\sum_{n=1}^{N} \sum_{k=1}^{3} \gamma_n(1,k)}$$

Example: Transition Probability

- $\sum_{n=1}^{N} \gamma_n(i,j)$ is the expected number of transitions between state s_i and s_j (given X and θ)
- we never take a hard decision on when the transition happened

Calculate Gamma (DP)

$$
\begin{aligned}
\gamma_n(i,j) &= P(z_n = s_j, z_{n-1} = s_i | X, \theta) \\
\text{(Bayes)} &= \frac{P(z_n = s_j, z_{n-1} = s_i, X | \theta)}{P(X|\theta)} \\
&= \frac{\alpha_{n-1}(i)\ a_{ij}\ \phi_j(x_n)\ \beta_n(j)}{\sum_{k=1}^{M} \alpha_N(k)}
\end{aligned}
$$



Baum-Welch: Properties

instance of Expectation Maximisation:

- iterative procedure
- guaranteed to convert to local maximum of the likelihood $P(X|\theta^{\text{new}})$
- sensitive to initialisation
- update formulae depend on state to output probability model $\phi_j(x_n)$